

## Constrained evolution of the sex comb in *Drosophila simulans*

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### Abstract

Male fitness is dependent on sexual traits that influence mate acquisition (pre-copulatory sexual selection) and paternity (post-copulatory sexual selection), and although many studies have documented the form of selection in one or the other of these arenas, fewer have done it for both. Nonetheless, it appears that the dominant form of sexual selection is directional, although theoretically, populations should converge on peaks in the fitness surface, where selection is stabilizing. Many factors, however, can prevent populations from reaching adaptive peaks. Genetic constraints can be important if they prevent the development of highest fitness phenotypes, as can the direction of selection if it reverses across episodes of selection. In this study, we examine the evidence that these processes influence the evolution of the multivariate sex comb morphology of male *Drosophila simulans*. To do this, we conduct a quantitative genetic study together with a multivariate selection analysis to infer how the genetic architecture and selection interact. We find abundant genetic variance and covariance in elements of the sex comb. However, there was little evidence for directional selection in either arena. Significant nonlinear selection was detected prior to copulation when males were mated to nonvirgin females, and post-copulation during sperm offence (again with males mated to nonvirgins). Thus, contrary to our predictions, the evolution of the *D. simulans* sex comb is limited neither by genetic constraints nor by antagonistic selection between pre- and post-copulatory arenas, but nonlinear selection on the multivariate phenotype may prevent sex combs from evolving to reach some fitness maximizing optima.

### Introduction

Male sexually selected traits typically evolve rapidly (Andersson, 1994; Arnqvist 1998) through both pre- and post-copulatory sexual selection (Partridge & Halliday, 1984). Pre-copulatory mechanisms of sexual selection include male–male competition and female mate choice, and post-copulatory mechanisms of sexual selection include sperm competition and cryptic female

choice (Parker, 1970; Eberhard, 1985; Andersson & Simmons, 2006; Hunt *et al.*, 2009). Given the complexity of the mechanisms of sexual selection, an understanding of the form and strength of selection that pre- and post-copulatory sexual selection impose is required to gain an understanding of the extravagance of the traits that they produce (Hunt *et al.*, 2009).

In the last decade, an increasing number of studies have used multivariate statistical techniques to describe the form and strength of selection on sexually selected traits (reviewed in Hunt *et al.*, 2009; Kingsolver & Diamond, 2011), and it is striking that directional selection is the dominant form of selection that has been documented (Hunt *et al.*, 2009; Kingsolver & Diamond, 2011). This is intriguing as, theoretically, populations

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should evolve towards areas of high fitness on fitness landscapes (Phillips & Arnold, 1989; Kingsolver & Diamond, 2011), and as populations move closer to these regions, selection should become stabilizing with moves in any direction acting to lower fitness (Chenoweth *et al.*, 2012). There are a number of mechanisms that may explain why populations never reach peaks on a fitness landscape, but one explanation is the presence of trade-offs that could arise either from the genetic covariance structure among traits under selection or from antagonism of selection on the multivariate phenotype across episodes of selection (e.g. pre- and post-copulatory episodes).

Genetic constraints may arise due to associations among traits (i.e. the genetic covariance structure), so selection on one will indirectly select on others (Cheverud, 1984; Phillips & Arnold, 1989; Blows & Brooks, 2003; Moore *et al.*, 2004; Bentsen *et al.*, 2006; Hunt *et al.*, 2007a; Pitcher *et al.*, 2014). If the genetic covariance or correlation ( $r_G$ ) between traits is negative with respect to each trait's (directional) effect on fitness (e.g.  $r_G < 0$  between two positively selected traits), this should limit selection towards an adaptive peak (Fear & Price, 1998; Blows & Hoffmann, 2005). Evidence consistent with bivariate genetic constraints has been found in a cricket (*Gryllus lineaticeps*; Wagner *et al.*, 2012), dung beetle (*Onthophagus taurus*; House & Simmons, 2005) and a cockroach (*Nauphoeta cinerea*; Moore *et al.*, 2004). More recently, however, a focus on bivariate correlations to infer constraints has been criticized, as the data from long-term studies suggest that populations do not evolve as predicted from bivariate genetic architecture alone (reviewed in Walsh & Blows, 2009). Instead, a multivariate approach that combines the genetic variance–covariance (**G**) matrix (i.e. the genetic variance across a suite of traits and the genetic covariances among them) with the vectors of linear selection gradients ( **$\beta$** ) (i.e. estimation of linear selection across suites of traits) has been advocated to assess the potential for genetic constraints (Walsh & Blows, 2009; Walling *et al.*, 2014).

If trade-offs occur between traits, they can also occur across discrete episodes of selection if trait values that increase fitness in one selective bout decrease it in another (Andersson & Simmons, 2006; Hunt *et al.*, 2009; Kingsolver & Diamond, 2011). For instance, if selection on a trait is positive during mate acquisition and negative during sperm competition, this can result in no net selection on traits (Hunt *et al.*, 2009). However, the empirical evidence for these sorts of trade-offs is mixed. For example, pre- and post-copulatory selection appear to be reinforcing in the guppy (*Poecilia reticulata*; Evans *et al.*, 2003), cricket (*Acheta domesticus*; Head *et al.*, 2006), fly (*Drosophila simulans*; Hosken *et al.*, 2008) and stalk-eyed fly (*Teleopsis dalmanni*; Rogers *et al.*, 2008). In contrast, episodes of pre- and post-copulatory selection are antagonistic in the water

strider (*Gerris lacustris*; Danielsson, 2001), dung beetles (*Onthophagus* species; Simmons & Emlen, 2006), firefly (*Photinus greeni*; Demary & Lewis, 2007), gulf pipefish (*Syngnathus scovelli*; Rose *et al.*, 2013) and the flour beetle (*Gnathocerus cornutus*; Okada *et al.*, 2014). So at least sometimes, the trait values that would be of highest fitness in one selective episode may not be highest in another selective bout, and therefore, evolution is constrained by antagonistic selection.

Many male *Drosophila* have a secondary sexual trait on their forelegs, the sex comb(s) (Kopp & True, 2002). These are used to grasp the female's abdomen and genitalia prior to and during copulation. The design of the sex combs is highly variable across closely related species, with comb and tooth number being especially variable (Markow *et al.*, 1996). Field and laboratory studies provide evidence that these interspecific patterns of phenotypic variation are partly due to sexual selection. For instance, during precopulatory sexual selection there is positive (directional) selection on comb size and comb symmetry in *D. bipectinata* (wild population; Polak *et al.*, 2004), whereas positive selection on tooth number has been reported in *D. melanogaster* (experimental lines; Promislow *et al.*, 1998). There is also post-copulatory selection on sex comb traits in *D. bipectinata*, with positive selection on comb size (artificial lines; Polak & Simmons, 2009) and nonlinear (disruptive) selection against intermediate tooth number in *D. melanogaster* (wild populations; Robinson *et al.*, 2012). However, a number of other studies have found less evidence for selection. For instance, no relationship between sex comb tooth number and mating success was found in either *D. melanogaster* (wild populations; Markow *et al.*, 1996; experimental lines; Snook *et al.*, 2013) or *D. pseudoobscura* (experimental lines; Snook *et al.*, 2013). This poses a paradox because although *Drosophila* sex combs have characteristics expected of a sexually selected trait (e.g. rapid divergence among lineages), the evidence that these characters are under strong sexual selection is inconsistent. One resolution may be that sex comb traits are the target of selection that has not been measured and/or selection on sex combs across pre- and post-copulatory selection is antagonistic.

In this study, we investigate the hypothesis that the evolution of the paired *D. simulans* sex comb is constrained by genetic constraints and/or antagonistic selection across episodes of sexual selection. Sexual selection has been intensely studied in *D. simulans* for a number of traits (e.g. Hosken *et al.*, 2008; Taylor *et al.*, 2008a,b; Ingleby *et al.*, 2014), and previous research suggests that sex comb tooth number is under negative directional selection through precopulatory mating success (Markow *et al.*, 1996). However, tooth number represents just one component of the multivariate comb phenotype and little is known about whether (and how) selection differs depending on whether it

occurs pre- vs. post-copulation. It is also unknown whether precopulatory selection is itself contingent on whether females have previously mated. Nonetheless, prior work has shown the single sex comb on the foretarsus of this species is functionally important, being used to grasp the female abdomen and genitalia and spread her wings prior to and during copulation (Sharma *et al.*, 2011). We therefore expect that overall comb morphology will be subject to directional selection. To start, we used a half-sib breeding design to estimate the genetic variance for and covariances among components of the sex comb (and body size). Next, we quantified the form and strength of sexual selection across four episodes of sexual selection: precopulatory selection when females were virgin or mated and post-copulatory sexual selection during sperm competition, when the focal male was first to mate (i.e. P1, sperm defence) or second to mate (i.e. P2, sperm offence).

## Materials and methods

### Fly stocks

Our laboratory wild-type populations of *D. simulans* were derived from 20 isolines (supplied by Centre for Environmental Stress and Adaptation Research, La Trobe University, Australia) that originally came from individuals that were caught in Tuncurry, eastern Australia, in March 2004. In the laboratory, these isolines were mixed and maintained for at least 7 years prior to the start of this study and have been found to be genetically and phenotypically variable for all traits that have been assayed (Hosken *et al.*, 2008; Wright *et al.*, 2009; Okada *et al.*, 2011; Sharma *et al.*, 2011). In addition to the wild-type population, laboratory populations of ebony flies, which carry a homozygous recessive phenotypic marker, were derived from a strain obtained from the Tucson stock centre and maintained as above for over 50 generations. The grey-black cuticle of ebony flies allows the easy discrimination between progeny of ebony females sired by ebony vs. wild-type males (Ashburner *et al.*, 2005). All population cages (wild-type and ebony) had an excess of 600 flies with overlapping generations and free mate choice. All stock and experimental offspring were maintained at 25 °C under a 12 : 12-h light:dark cycle and maintained on *Drosophila* culture medium (Jazz Mix *Drosophila* Food, Fisher Scientific; and *Drosophila* Quick Mix Medium, Blades Biological) with an excess of food. This reduces the risk of environmental influences affecting mating and remating probabilities because of stress response (Zera & Harshman, 2001).

### Breeding design

#### Parental generation

For our experimental breeding design, wild-type flies were initially collected from population cages. Egg

laying vials were placed in the cages of two wild-type populations daily and left for 24 h. These vials were incubated until peak eclosion (ca. 8–9 days after egg laying). Offspring that eclosed overnight were killed and virgins were collected ca. 7 h later (Sharma *et al.*, 2010). Virgin males were maintained in standard culture vials, with ca. 80 males per vial. Virgin females were aspirated into ca. 800 individual vials containing culture medium. These virgin females and males were the parents for our design and were 3 days old before breeding commenced to ensure full sexual receptivity (Manning, 1967).

#### Breeding and rearing

A conventional half-sibling breeding design was used (Lynch & Walsh, 1998), where 130 sires were each mated with five dams. Details of the mating regime are as follows: a sire was housed with a randomly selected, virgin female for 24 h to maximize the probability that the pair would mate. The following day, the male was aspirated from the vial and transferred to a new vial that contained a virgin female for 24 h. The process was repeated three more times until the sire had been housed with a total of five dams. The mated dams were housed singly in oviposition vials and transferred daily to new oviposition vials for a total of 4 days. The oviposition vials were stored at 25 °C for 12 days under a 12/12-h light:dark cycle until the offspring began to emerge. Six days after the first eclosion, the offspring were collected, labelled and frozen at –20 °C for subsequent dissection, measurement and quantitative genetic analysis.

### Multivariate sexual selection

#### Experimental design

For experimental mating assays, a sample of ebony and wild-type flies (not the same as those that were used for the breeding design) were collected as virgins from population cages using the protocols described above ('Parental generation'). Virgin females and males were used for mating trials when the females were 3 days old and males were 3–4 days old, to ensure full sexual receptivity (Manning, 1967). Mating trials began at the beginning of the photophase of the light:dark cycle as this is when the flies are most reproductively active (Sakai & Ishida, 2001). In all trials, each male was aspirated into a female housing vial and continuously observed for 2 h during which courtship (i.e. wing flicking, wing vibration, leg rubbing and licking) and mating were recorded (Spieth, 1974).

#### Sex comb morphology and precopulatory sexual selection

In the first part of the study, we investigated whether variation in sex comb morphology predicts mating success with virgin females (Virgin Trial) or with mated

females (Nonvirgin Trial). To do this, we used no-choice mating assays that are a standard method to assess overall male attractiveness (e.g. Hedge & Krishna, 1997; Koref-Santibanez, 2001; Gowaty *et al.*, 2002; Yenissetti & Hedge, 2003; Shackleton *et al.*, 2005), and the results of assays with single males and multimales are the same (Taylor *et al.*, 2008a,b). During Virgin Trials, males that courted but were rejected ( $n = 154$ ) or courted and mated ( $n = 340$ , total  $n = 494$ ) were separated from the females and frozen at  $-20\text{ }^{\circ}\text{C}$  for morphometric measurement. During Nonvirgin Trials, we used a new set of flies that were derived from the same stock population. The females were once mated, but detailed observation of their mating behaviour was not recorded. All females were 7 days old, having mated 4 days before their second exposure to virgin males. The mating procedure in this trial was identical to that described above (Virgin Trial). All males that courted but were rejected ( $n = 329$ ) or courted and mated ( $n = 154$ , total  $n = 483$ ) were frozen at  $-20\text{ }^{\circ}\text{C}$  for morphometric measurement.

#### *Sex comb morphology and post-copulatory sexual selection*

In the second part of the selection study, we investigated whether variation in sex comb morphology predicts fertilization success. Ebony females were sequentially mated with a focal, wild-type male followed by an ebony male (paternity defence – P1) or an ebony male followed by a focal, wild-type male (paternity offence – P2). Males mated once only and in a single role – defensive or offensive. During the observation period, if copulation occurred, the male was removed from the chamber, aspirated into an Eppendorf tube and stored at  $-20\text{ }^{\circ}\text{C}$  for dissection and measurement. Following the first mating, females were transferred daily into fresh food vials to oviposit for 4 days before their second exposure to virgin males. The second mating procedure for mated females was identical to that described above. Ebony females that did not mate with the second mating partner during the 2-h assay were excluded from the data set, along with their first mate ( $n \sim 600$  – *D. simulans* are reluctant to mate, particularly with mutant strains). Following their second mating, twice-mated females were once again transferred daily into fresh food vials to oviposit for 4 days. On the 5th day, the female was aspirated into an Eppendorf tube and stored at  $-20\text{ }^{\circ}\text{C}$ . Vials that had contained the mated females were stored at  $25\text{ }^{\circ}\text{C}$  and monitored daily until offspring emerged. Seven days after the first emergence, the vials were inverted and stored in the freezer and the ebony and wild-type offspring from each of the female's eight vials was subsequently counted to determine the number of offspring that were sired by the focal (i.e. wild-type) male during defensive (P1,  $n = 308$ ) or offensive mating (P2,  $n = 355$ ).

#### *Dissection and morphometric measurement*

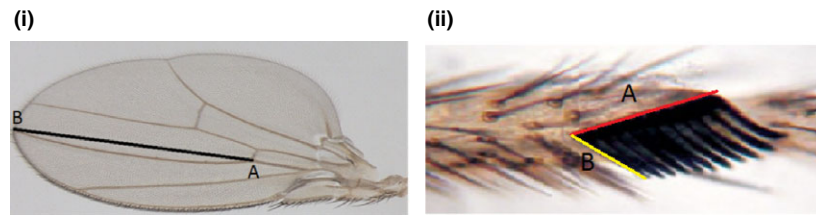
The left and right forelegs and wings of focal, wild-type males or sons from our breeding design were carefully pulled free from the body of each male and then mounted on glass slides in a droplet of Hoyer's medium. Digital images for wings ( $\times 30$ ) and sex combs ( $\times 100$ ) were captured using a Leica dissecting microscope (M125) connected to a Leica camera (DFC295). Wing length and sex comb components were measured using ImageJ v1.46r (RSB National Institute of Mental Health, USA) (Fig. 1).

We used wing length (WL) as an index of body size (Markow & Ricker, 1992; Gilchrist & Partridge, 1999; Sharma *et al.*, 2011), and both left and right wings of each male were measured and an average value was calculated. Three components of sex comb morphology were measured: the comb length (CL); tooth length (TL), measured as the average length of the first, third and fifth teeth; and comb tooth number (TN) (Fig. 1). All sex comb characteristics including CL, TL and TN were estimated as the average of the measurements on the left and right body sides. The precision of the measurements was assessed by blindly measuring all traits twice on a subsample of wings and sex combs ( $N = 20$ ). Two measures of the same trait were tightly correlated (TL:  $r^2 = 0.919$ ,  $P < 0.05$ ; CL:  $r^2 = 0.982$ ,  $P < 0.001$ ; TN:  $r^2 = 1.00$ ,  $P < 0.001$ ; WL:  $r^2 = 0.992$ ,  $P < 0.001$ ).

#### **Statistical analysis**

##### *Genetic analyses*

Data were analysed using animal models fitted with restricted maximum likelihood in ASReml (version 3.0; VSN International Ltd) with assumed Gaussian errors (see Wilson *et al.*, 2010). First, we tested for additive genetic variance using univariate models fitted to each of the sex comb component traits (comb length CL, tooth length TL and tooth number TN) and size (wing length WL). Each model contained the mean as a fixed effect and random effects of additive genetic merit and a 'maternal identity' effect. The latter was included to protect against upward bias from maternal (or other common environment) effects shared by full-sibs. For each trait, we compared this to a reduced model with the additive effect dropped using a likelihood ratio test and assuming that twice the difference in log-likelihoods is distributed as a 50 : 50 mix of  $\chi^2_1$  and  $\chi^2_0$  (subsequently denoted  $\chi^2_{0,1}$ ). Having detected significant genetic variance in all traits (see results), we formulated a multivariate animal model which was used to estimate the additive variance–covariance matrix (**G**) and derived parameters. To facilitate convergence in the multivariate model, traits were scaled to unit variance by dividing by their (observed) standard deviations. Heritability ( $h^2$ ) was estimated for each trait as  $V_A/V_P$ , where  $V_A$  is the additive genetic variance and  $V_P$  is the phenotypic variance, determined as the sum of  $V_A$ ,  $V_M$



**Fig. 1** Morphological measures of male *Drosophila simulans*: (i) wing and (ii) sex comb. The length of the wing was measured as the distance between points a and b. Three components of the sex comb were measured: comb length (CL; a); tooth length (TL), which was measured as the average length of the first (b), third and fifth teeth; and comb tooth number (TN).

(maternal variance) and  $V_R$  (residual variance). We similarly estimated the magnitude of the maternal effect as  $m^2$ , where  $m^2 = V_M/V_P$ . Genetic correlations ( $r_G$ ) were determined for each pair of traits (1,2) as  $r_{G(1,2)} = \text{COV}_{A(1,2)} / (V_{A1} * V_{A2})^{0.5}$ , where  $\text{COV}_A$  is the estimated additive genetic covariance. For comparison, we also estimated the corresponding phenotypic correlations  $r_P$ .

#### Multivariate selection analysis

To determine whether male phenotypic traits (CL, TL, TN and WL) influenced fitness during precopulatory or post-copulatory selection, we used a standard multivariate selection analysis approach. In precopulatory bouts of selections, a male was assigned a score of 1 if the male courted and mated and a 0 if the male courted only. In these mating success trials, the female was always presented with a wild-type male to increase the likelihood that a male would attempt to court and mate. As a consequence, we would have been unable to determine the number of offspring that were sired by the focal male when mating a previously mated female without extensive genotyping work, hence the binary fitness measure. In post-copulatory, fertilization success trials, male fitness was assigned a continuous value – the number of offspring that were sired by the focal male which ranged from 0 to 200. The mating and fertilization success response variables were transformed to relative fitness by dividing individual scores by the mean for each data set. The male phenotypic traits were standardized to zero means and unit variances as suggested by Lande & Arnold (1983). We then fitted a separate linear multiple regression for each of the four bouts of selection to estimate linear selection gradients when females were virgins ( $\beta_V$ ) or previously mated ( $\beta_M$ ) or the focal male mated in a defensive role ( $\beta_{P1}$ ) or an offensive role ( $\beta_{P2}$ ) (Lande & Arnold, 1983). Next, we applied a quadratic regression model including all linear, quadratic and cross-product (i.e. correlational) terms to estimate the matrix of nonlinear selection gradients for males when females were virgin ( $\gamma_V$ ) or previously mated ( $\gamma_M$ ) or the focal male mated in a defensive role ( $\gamma_{P1}$ ) or an offensive role ( $\gamma_{P2}$ ). Quadratic regression

coefficients were doubled to yield the standardized nonlinear selection gradients (see Stinchcombe *et al.*, 2008). As our binary and continuous fitness measures did not conform to a normal distribution, we used a resampling procedure to assess the significance of our linear and nonlinear selection gradients. Our fitness scores were randomly shuffled across individual phenotypes 10 000 times to generate a null distribution of pseudoselection gradients expected in the absence of a causal phenotype–fitness relationship (Mitchell-Olds & Shaw, 1987). The probability that the gradient pseudo-estimate was equal to or less than the original estimated gradient (out of 9999 permutations) was then tested. We conducted separate randomization analyses for the multiple regression models for directional selection (i.e. model containing only linear terms) and for the full quadratic model (i.e. model containing linear, quadratic and correlational terms).

To establish the extent of nonlinear selection acting on male phenotypic traits, we conducted a canonical analysis using the approach suggested by Reynolds *et al.* (2010). The analysis generates a new matrix that consists of vectors of linear selection described by theta ( $\theta_i$ ) and nonlinear selection that are described by eigenvalues ( $\lambda_i$ ) and their corresponding eigenvectors ( $\mathbf{m}_i$ ). Tests of the significance of the eigenvalues were conducted using the permutation procedure outlined in Reynolds *et al.* (2010). We used thin-plate splines (Green & Silverman, 1994) to visualize the major axes of the fitness surfaces extracted from the canonical rotation of  $\gamma_M$ , and  $\gamma_{P2}$ . Tps functions in the fields package of R (version 2.13.0; available via <http://www.r-project.org>) were used to fit spline surfaces using the value of the smoothing parameter ( $\lambda$ ) that minimized the generalized cross-validation (GCV) score. We then plotted surfaces in R using both the perspective and contour map views. Finally, to test whether the linear, quadratic and correlational selection gradients differed when females had previously mated compared to when males mated in the offensive role, we used a sequential model building approach (partial F-test) (Draper & John, 1988; see Chenoweth & Blows (2005) for a detailed description of this procedure).

## Results

### Genetic architecture

Comparison of full and reduced univariate models indicated significant additive genetic variance for comb length (CL:  $\chi^2_{0,1} = 25.0$ ,  $P < 0.001$ ), tooth length (TL:  $\chi^2_{0,1} = 5.48$ ,  $P = 0.010$ ), tooth number (TN:  $\chi^2_{0,1} = 40.2$ ,  $P < 0.001$ ) and wing length (WL:  $\chi^2_{0,1} = 4.78$ ,  $P = 0.014$ ). Estimates of maternal variance were non-zero in all cases except for TN where  $V_M$  was bound at zero (full results not shown), so we formulated the multivariate model with a  $4 \times 4$   $\mathbf{G}$  matrix but a  $3 \times 3$  maternal effect covariance matrix (i.e. no maternal effect on TN). Under this multivariate model,  $h^2$  estimates for sex comb components ranged from moderate to high (Table 1). The heritability of wing length (which is a proxy for body size) was similar to previously published heritability of body size for *Drosophila* ( $h^2 \sim 0.4$ ; Robertson, 1957;  $\sim 0.5$ ; Coyne & Beecham, 1987). All genetic correlations between sex comb component traits were positive and nominally significant (based on  $|r_G| > 2SEs$ ; Table 2). Genetic correlations between wing length and all sex comb components were also positive although not significantly for WL and TL. While noting that estimated standard errors are approximate and so not necessarily robust for formal

**Table 1** Phenotypic means and estimates of heritability ( $h^2$ ) and maternal effect ( $m^2$ ) for male body size and sex comb components ( $N$  sires = 110,  $N$  offspring = 1449). Estimates are from the multivariate animal model (see text for details).

Trait (unit)	Mean	$h^2$ (SE)	$m^2$ (SE)
Comb length ( $\mu\text{m}$ )	58.23 $\pm$ 0.12	0.61 $\pm$ 0.08	0.07 $\pm$ 0.02
Tooth length ( $\mu\text{m}$ )	39.23 $\pm$ 0.05	0.26 $\pm$ 0.11	0.27 $\pm$ 0.05
Tooth number	9.90 $\pm$ 0.02	0.53 $\pm$ 0.06	NA
Wing length (mm)	1154.08 $\pm$ 1.12	0.45 $\pm$ 0.06	0.40 $\pm$ 0.06

**Table 2** Additive genetic correlations above the diagonal and phenotypic correlations below the diagonal for sex comb components: comb length (CL), tooth length (TL), tooth number (TN) and wing length (WL). Significant genetic ( $|r_G| > 2SEs$ ) and phenotypic correlations are in bold (after Bonferroni correction).

	CL	TL	TN	WL
Comb length (CL)		<b>0.31 <math>\pm</math> 0.15</b>	<b>0.89 <math>\pm</math> 0.03</b>	<b>0.64 <math>\pm</math> 0.10</b>
Tooth length (TL)	<b>0.27 <math>\pm</math> 0.02</b>		<b>0.89 <math>\pm</math> 0.03</b>	0.27 $\pm$ 0.23
Tooth number (TN)	<b>0.84 <math>\pm</math> 0.01</b>	0.01 $\pm$ 0.03		<b>0.70 <math>\pm</math> 0.11</b>
Wing length (WL)	<b>0.45 <math>\pm</math> 0.02</b>	<b>0.44 <math>\pm</math> 0.02</b>	<b>0.23 <math>\pm</math> 0.02</b>	

inference, the model was a significantly better fit to the data than a reduced version in which all off-diagonal (i.e.  $\text{COV}_A$ ) terms in the  $\mathbf{G}$  matrix were constrained to zero ( $\chi^2_6 = 112$ ,  $P < 0.001$ ). Thus, it is clear that  $\mathbf{G}$  contains significant additive genetic covariance among the traits, and estimates are uniformly positive across all trait pairs.

### Sexual selection on sex combs

Rather surprisingly, given the evidence from a previous study in *D. simulans* (Markow *et al.*, 1996) we found no evidence of significant directional selection (i.e.  $\beta$  – linear selection that increases/decreases the trait mean) acting on any component of the sex comb in any of the four selective contexts (Table 3). However, we found evidence for nonlinear selection, which acted differently in each context. There are three different forms of nonlinear selection (i.e.  $\gamma$  coefficients that describe the curvature of nonlinear selection on individual traits): (a) stabilizing where  $\gamma$  coefficients are negative and individuals with intermediate trait values have highest fitness, (b) disruptive where  $\gamma$  coefficients are positive and individuals with extreme low or high trait values have highest fitness and (c) correlational selection where pairs of traits are jointly acted upon (Hunt *et al.*, 2009). We find evidence for all three forms of nonlinear selection.

### Precopulatory sexual selection

Nonlinear selection was weak and nonsignificant when males courted virgin females with the exception of significant positive correlational selection between tooth number (TN) and wing length (WL) (Table 3A). Canonical rotation of the  $\gamma$  matrix of nonlinear selection gradients produced one positive and three negative eigenvalues, which describe the curvature of selection on the major axes of selection, rather than on individual traits (Table 4A; i.e. positive eigenvalue is indicative of disruptive selection along  $\mathbf{m}_1$  and negative eigenvalue is indicative of stabilizing selection along  $\mathbf{m}_2$ – $\mathbf{m}_4$ ). However, selection on the eigenvectors ( $\mathbf{m}_1$ – $\mathbf{m}_4$ ) was nonsignificant (Table 4A).

Nonlinear selection was stronger when males courted nonvirgin females. There was significant stabilizing (negative  $\gamma$ ) selection on tooth length (TL) and disruptive (positive  $\gamma$ ) selection on the tooth number (TN) as well as positive correlational selection between tooth length (TL) and wing length (WL) (Table 3B). Canonical rotation of the  $\gamma$  matrix of nonlinear selection gradients produced a combination of disruptive selection along the  $\mathbf{m}_1$  and  $\mathbf{m}_2$  axis and stabilizing selection along the  $\mathbf{m}_3$  and  $\mathbf{m}_4$  axis; however, there was only significant selection along eigenvector  $\mathbf{m}_4$  (Table 4B). This axis of significant selection for the nonvirgin mating phase shows stabilizing (negative  $\gamma$ ) selection which

**Table 3** The vector of standardized linear selection gradients ( $\beta$ ) and the matrix of standardized nonlinear gradients ( $\gamma$ ) for sex comb morphological traits in male *Drosophila simulans* during precopulatory sexual selection when a male courted and/or mated a (A) virgin female or (B) nonvirgin female and during post-copulatory selection in a competitive role when a male mated in a (C) defensive role (i.e. P1) or (D) offensive role (i.e. P2).

$\beta$	$\gamma$			
	CL	TL	TN	WL
A. Standardized selection gradients when a male courted and/or mated a virgin female				
CL	0.020	0.264		
TL	0.019	-0.051	-0.012	
TN	-0.029	-0.161	-0.041	-0.012
WL	-0.004	-0.108	0.027	0.196*
B. Standardized selection gradients when a male courted and/or mated a nonvirgin female				
CL	-0.172	0.104		
TL	-0.051	0.031	-0.300*	
TN	0.010	-0.249	-0.022	0.422*
WL	0.121	0.069	0.228**	-0.017
C. Standardized selection gradients when a male mated in a defensive role (P1)				
CL	-0.108	-0.030		
TL	-0.010	0.099	-0.106	
TN	0.127	-0.118	0.041	0.228
WL	-0.119	-0.018	-0.073	-0.121
D. Standardized selection gradients when a male mated in an offensive role (P2)				
CL	-0.055	0.806**		
TL	0.038	0.025	-0.154	
TN	0.094	-0.509*	-0.026	0.25
WL	-0.047	-0.282*	0.119	0.157

CL, comb length; TL, tooth length; TN, tooth number; WL, wing length. Randomization tests: \* $P < 0.05$ , \*\* $P < 0.01$ , \*\*\* $P < 0.001$ .

†Nonlinear selection gradients include quadratic ( $z_{ii}^2$ ) gradients on the diagonal and correlational ( $z_{ij}$ ) gradients below the diagonal.

we visualized with  $\mathbf{m}_1$  that had the largest, albeit non-significant disruptive (positive  $\gamma$ ) eigenvalue. These represent parts of the fitness surface that curve downwards and upwards respectively to create a saddle-like fitness surface in the  $\mathbf{m}_1$ - $\mathbf{m}_4$  plot (Fig. 2a). Along the  $\mathbf{m}_4$  axis, highest fitness occurred along a ridge which corresponds with intermediate values and was heavily influenced by tooth length (TL) and wing length (WL) (i.e. in each row of M in Table 4, the magnitude of the values indicates the contribution of individual traits to an eigenvector). A contour-view visualization of the same fitness surface with an overlay of the data points shows that many of the males are spread along the ridge on the  $\mathbf{m}_4$  axis (Fig. 2b).

### Post-copulatory sexual selection

Nonlinear selection was weak and nonsignificant when males mated in a defensive role with the exception of

**Table 4** Linear ( $\theta_i$ ) and nonlinear ( $\lambda_i$ , the eigenvalue) selection gradients and the M matrix† of eigenvectors ( $\mathbf{m}_i$ ) from the canonical analysis of  $\gamma$  for (A) virgin mating success, (B) nonvirgin mating success, (C) P1 experiment and (D) P2 experiment.

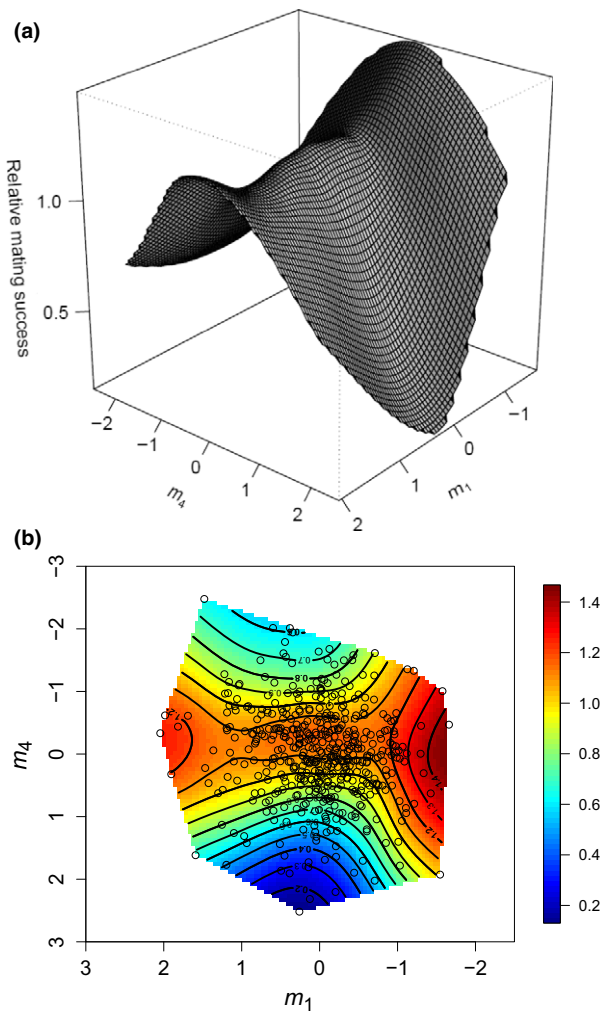
	$\theta_i$	$\lambda_i$	M			
			CL	TL	TN	WL
A. Canonical analysis of virgin mating success						
$\mathbf{m}_1$	0.029	0.286	0.696	0.428	-0.575	-0.038
$\mathbf{m}_2$	0.015	-0.010	0.629	-0.148	0.623	0.441
$\mathbf{m}_3$	-0.020	-0.139	-0.021	-0.629	-0.530	0.568
$\mathbf{m}_4$	-0.011	-0.249	0.345	-0.632	-0.007	-0.693
B. Canonical analysis of nonvirgin mating success						
$\mathbf{m}_1$	0.086	0.565	<b>-0.485</b>	-0.059	<b>0.869</b>	-0.074
$\mathbf{m}_2$	0.076	0.004	-0.674	-0.379	-0.440	-0.456
$\mathbf{m}_3$	-0.132	-0.099	0.555	-0.573	0.223	-0.560
$\mathbf{m}_4$	-0.128	-0.515**	0.047	<b>0.724</b>	0.016	<b>-0.687</b>
C. Canonical analysis of P1						
$\mathbf{m}_1$	0.184	0.374	-0.145	0.133	0.715	-0.671
$\mathbf{m}_2$	0.049	0.173	-0.489	-0.258	0.542	0.633
$\mathbf{m}_3$	-0.070	-0.022	-0.614	-0.575	-0.379	-0.385
$\mathbf{m}_4$	-0.023	-0.200	0.602	-0.764	0.227	-0.041
D. Canonical analysis of P2						
$\mathbf{m}_1$	0.078	1.204*	<b>-0.824</b>	0.0003	<b>0.487</b>	0.288
$\mathbf{m}_2$	0.052	0.096*	-0.182	<b>-0.444</b>	0.197	<b>-0.855</b>
$\mathbf{m}_3$	0.047	-0.052	0.527	0.008	0.846	0.078
$\mathbf{m}_4$	-0.068	-0.216	0.095	-0.896	-0.090	0.424

Randomization tests: \* $P < 0.05$ , \*\* $P < 0.01$ , \*\*\* $P < 0.001$ .

†Values in bold, contributed most to that eigenvector ( $\mathbf{m}_i$ ).

significant disruptive selection (positive  $\gamma$ ) on wing length (Table 3C). Canonical rotation of the  $\gamma$  matrix of quadratic selection gradients produced a combination of disruptive selection along the  $\mathbf{m}_1$  and  $\mathbf{m}_2$  axis and stabilizing selection along the  $\mathbf{m}_3$  and  $\mathbf{m}_4$  axis; however, selection along these vectors ( $\mathbf{m}_1$ - $\mathbf{m}_4$ ) was nonsignificant (Table 4C).

Nonlinear selection was stronger during competitive mating when males mated in the offensive role (P2). There was disruptive (positive  $\gamma$ ) selection on comb length (CL) and negative correlational selection between comb length (CL) and tooth number (TN) and comb length (CL) and wing length (WL) (Table 3D). Canonical rotation of the  $\gamma$  matrix of quadratic selection gradients produced a combination of disruptive selection along the  $\mathbf{m}_1$  and  $\mathbf{m}_2$  axis and stabilizing selection along the  $\mathbf{m}_3$  and  $\mathbf{m}_4$  axis, but selection along these vectors was only significant for  $\mathbf{m}_1$  and  $\mathbf{m}_2$ . These axes of significant selection for the competitive, offensive mating phase (P2) showed disruptive selection along the  $\mathbf{m}_1$  and  $\mathbf{m}_2$  axes which curved the fitness upwards to create an inverted fitness surface in the  $\mathbf{m}_1$ - $\mathbf{m}_2$  plot (Fig. 3a). Along the ridge of highest fitness (i.e. intermediate values of  $\mathbf{m}_1$  and positive values of  $\mathbf{m}_2$ ), high paternity was correlated with a long sex comb, few but long comb teeth and large body size. However, a contour-view visualization of the same fitness surface with

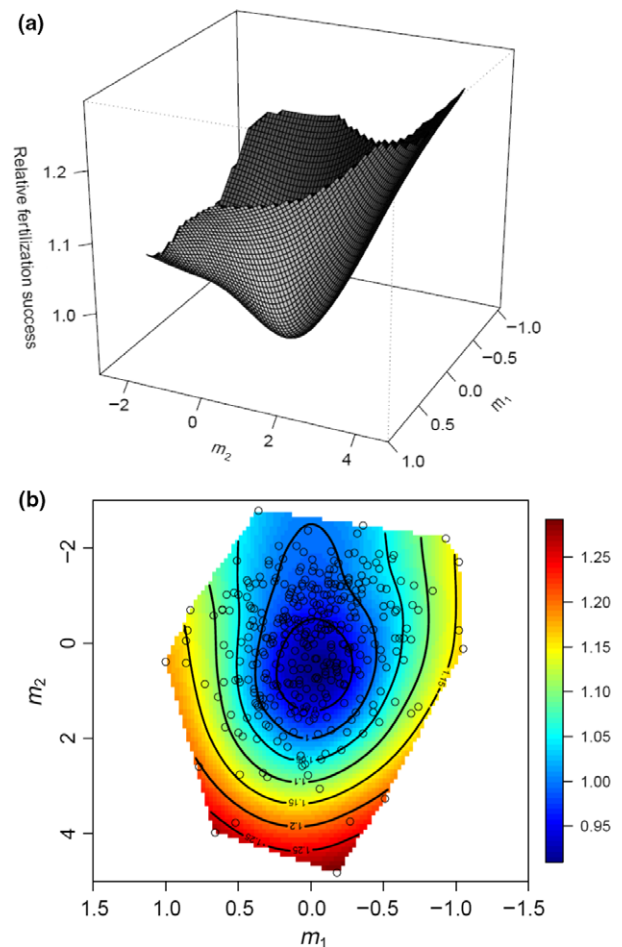


**Fig. 2** Thin-plate spline visualizations (a: perspective view; and b: contour view) of the two major axes of nonlinear selection ( $\mathbf{m}_1$  and  $\mathbf{m}_4$ ) on the fitness surface when males courted nonvirgin females. In the contour view, red-to-orange coloration represents regions of highest fitness, whereas blue coloration represents regions of lowest fitness. Individual data points are provided as black circles on the surface.

an overlay of the data points shows that few males occupy this region on the landscape (Fig. 3b).

### The strength and form of linear and nonlinear selection across episodes

To test for possible differences in selection on the sex comb and body size (i.e. WL) during bouts of significant pre- and post-copulatory sexual selection, we compared the strength of linear, quadratic and correlational selection across selective bouts. The strength of linear ( $F_{4,826} = 1.192$ ,  $P = 0.313$ ), quadratic ( $F_{4,818} = 1.576$ ,  $P = 0.179$ ) and correlational selection ( $F_{6,806} = 0.469$ ,



**Fig. 3** Thin-plate spline visualizations (a: perspective view; and b: contour view) of the two major axes of nonlinear selection ( $\mathbf{m}_1$  and  $\mathbf{m}_2$ ) on the fitness surface when males mated in the offensive role (P2). In the contour view, red-to-orange coloration represents regions of highest fitness, whereas blue coloration represents regions of lowest fitness. Individual data points are provided as black circles on the surface.

$P = 0.759$ ) did not differ significantly between these bouts of selection.

### Discussion

We find that there is substantial genetic variation in the male sex comb trait components which are positively genetically correlated with each other and with body size. However, there was no evidence of directional selection on the sex comb across any bout of sexual selection. Thus, contrary to our predictions, it is the absence of directional selection that is the primary limitation to the evolution of the *D. simulans* sex comb rather than genetic constraints arising from among-trait covariance and/or antagonistic linear selection across



episodes of selection. Whilst evidence for linear selection was conspicuous by its absence, we did find complex patterns of significant nonlinear selection. In particular, we found disruptive selection acting on male sex combs during post-copulatory selection when females are already mated.

### Genetic (co)variance among components of the sex comb

A breadth of studies find that sexually selected traits harbour abundant genetic variation (reviewed in Roff & Mousseau, 1987; Houle, 1992; Pomiankowski & Moller, 1995; Walsh & Blows, 2009). Our average  $h^2$  estimate for sex comb components ( $h^2 = 0.46$ ) is high and comparable with other  $h^2$  estimates for morphological traits (Roff & Mousseau, 1987; Houle, 1992; Pomiankowski & Moller, 1995). The maintenance of genetic variation in sexually selected traits is an evolutionary puzzle, and a number of models have been developed to explain the phenomena (Taylor & Williams, 1982; Mousseau & Roff, 1987; Pomiankowski & Moller, 1995; Rowe & Houle, 1996). Here, it appears that the lack of significant directional selection coupled with stabilizing and disruptive selection (which may promote genetic variation) has maintained genetic variance in the sex comb. We also found positive genetic correlations between component traits of the comb and body size which should result in positively correlated indirect selection responses.

### Linear selection on the sex comb across selective episodes

A previous study of *D. simulans* found that directional selection during precopulatory sexual selection favoured fewer teeth in the comb, whereas we found no evidence that directional selection acts on sex comb components during any bout of selection. More generally, the evidence that selection acts on components of the sex comb of *Drosophila* species is mixed. In part, this may be a result of experimental design – typically, estimates of selection on the sex comb are univariate (Markow *et al.*, 1996; Promislow *et al.*, 1998; Polak *et al.*, 2004; Polak & Simmons, 2009; Snook *et al.*, 2013) even though this may underestimate the strength of selection (Blows & Brooks, 2003). For instance, if nonlinear selection was acting, it could result in linear selection gradients being estimated that simply cross two points of a nonlinear selection gradient (Hunt *et al.*, 2009). The results of this study, and from a field study of *D. melanogaster*, where sexual selection on the sex comb was disruptive (Robinson *et al.*, 2012), suggest that this may be an oversight as nonlinear selection was the dominant form of selection.

Given that directional selection on male sex comb components was absent, it is clear that trade-offs

(between component traits and/or pre- vs. post-selective episodes) are neither present nor required to explain evolutionary stasis. Among previous studies of sexual selection on male traits, precopulatory selection for elaborate male traits is often reinforced by post-copulatory fertility benefits (Rogers *et al.*, 2008), sons with high fertilization success (Hosken *et al.*, 2008) and/or high-quality sons (Head *et al.*, 2006). In *D. simulans*, precopulatory selection acting on the sex comb is weak, so it seems unlikely that females exercise mate choice on the basis of male sex comb morphology. Furthermore, during post-copulatory sexual selection more than one sex comb phenotype is correlated with fertilization success during competitive mating, and therefore, it seems unlikely that sex combs provide a clear signal of sire or offspring reproductive quality.

### Nonlinear selection on the sex comb across selective episodes

When females were already mated, precopulatory sexual selection favours males with intermediate tooth length (TL) and wing sizes which resulted from a blend of stabilizing and correlational selection on these traits. A similar pattern of stabilizing selection has been found in *D. melanogaster* following successful (artificial) linear selection for high or low tooth number (Ahuja & Singh, 2008). After ten generations of relaxed selection, tooth number regressed back to intermediate, control numbers, demonstrating the action of net stabilizing selection on this component of the comb (Ahuja & Singh, 2008). Interestingly, males within the low tooth number lines were less likely to successfully mate if they had very few teeth compared to those that had more sex comb teeth. However, among the control and high tooth number lines, the effect of tooth number on mating success was nonsignificant (Ahuja & Singh, 2008).

Here, nonlinear post-copulatory selection on the sex combs was stronger and disruptive when measured as sperm offence. As the significant eigenvalues ( $\lambda$ ) are positive, it suggests that the fitness surface is concave and best described as a bowl (Fig. 3a; Hunt *et al.*, 2009), and along the height of the bowl, fertilization success is approximately equivalent (Fig. 3b). Two other studies have shown that particular morphologies of *Drosophila* sex combs enhance competitive fertilization success. In *D. bipunctinata*, artificial selection was used to develop lines with relatively short or long combs, and relatively long combs were found to confer an advantage during sperm offence (Polak & Simmons, 2009). In contrast, in a field study of *D. melanogaster*, sexual selection on the sex comb was disruptive (Robinson *et al.*, 2012) as we report here.

The patterns of selection that we found may be explained if extreme combinations of sex comb components are most effective at grasping the female and

aligning her genitalia during insemination, with intermediate combinations being less effective. Similar patterns of disruptive selection have been found in naturally selected traits (Brodie, 1992; Smith, 1993; Bolnick, 2004) and sexually selected traits (Blows *et al.*, 2003), and in three of these studies, competition for limiting resources appears to generate this pattern of selection (Smith, 1993; Blows *et al.*, 2003; Bolnick, 2004). For instance, the African finch exhibit small or large bill size and feed exclusively on soft- or hard-seeded sedge, respectively (Smith, 1993). In the three-spine stickleback, intraspecific competition selects for extreme trophic morphology (i.e. large or small gill raker length) (Bolnick, 2004), and female choice selects for rare male phenotypes in guppies (Blows *et al.*, 2003). The wider implication of this pattern of selection is subject to debate but theoretically can force niche expansion (Roughgarden, 1972), sexual dimorphism (Slatkin, 1984; Bolnick & Doebeli, 2003) and speciation (Doebeli, 1996; Dieckmann & Doebeli, 1999).

### Nonlinear selection and genetic correlations between sex comb components

In this study, all genetic correlations between body size and the sex comb components were positive so that genotypes predisposing to larger size also result in longer combs with both more numerous and longer teeth. However, highest post-copulatory, paternity offence (i.e. P2) was correlated with a long sex comb but few comb teeth (i.e. intermediate  $m_1$  and positive  $m_2$ ), yet the positive genetic covariance between these sex comb traits means that few male genotypes occupy this region of the landscape. This may reflect an underlying mechanistic constraint as sex combs are positively allometric (Sharma *et al.*, 2011), so the scaling of sex comb trait components with body size largely prevents this combination. Evidence from other species suggests that the cause of genetic covariance may originate from developmental or functional constraints that place limits on trait combinations. For example, negative genetic covariance between the call rate and chirp duration of a cricket (Wagner *et al.*, 2012) and ejaculate size and sperm quality in a cockroach (Moore *et al.*, 2004) may reflect the energetics of calling (Wagner *et al.*, 2012) and sperm production (Moore *et al.*, 2004), whereas the negative covariance among colour pattern components in a guppy (Brooks & Endler, 2001) and the correlated evolution of beak morphology and vocal repertoire of Darwin's finches (Podos, 2001) may be due to physical constraints. For instance, in the guppy, spots occupied by one colour may preclude another (Brooks & Endler, 2001) and, in finches, beaks that become adapted for increased bite force are less able to perform rapid movements that are required for certain songs (Podos, 2001).

### Opportunity for sexual selection across selective episodes

*Drosophila simulans* belong to a clade in which female remating is infrequent and females can be more choosy after mating as they can use stored sperm to continue to produce offspring (Taylor *et al.*, 2007, 2008a,b). More broadly in *Drosophila sp.*, it is striking that secondary sexual traits, such as the sex comb, are only present in clades where females rarely remate. Theoretically, this should increase the variance in male mating success and thus the opportunity for selection on male secondary sexual characters (Emlen & Oring, 1977; Markow, 2002; Collet *et al.*, 2012). Our results are partially consistent with this expectation, with no evidence of selection on male sex combs during any bout of selection when females are virgin, whereas we detect some nonlinear selection on sex combs when females are mated. For example, during precopulatory selection, virgin females rejected ~30% of male courtship displays and this increased to ~60% when females were mated which was matched by the detection of selection on the sex comb, albeit weak selection. Similarly, post-copulatory sexual selection on sex combs was only significant for sperm offence – so selection on the sex combs is acting through males' ability to displace a mated female's stored sperm.

### Conclusions

Given the lack of directional selection acting on the sex comb, formal analysis of the extent to which the covariance structure in  $\mathbf{G}$  constrains a selection response is not particularly informative. Studies that have implemented this approach show that genetic covariances constrain evolution sometimes (Hine *et al.*, 2004; Hunt *et al.*, 2007a,b; Von Homrigh *et al.*, 2007; Hall *et al.*, 2010; Ingleby *et al.*, 2014) but not always (Ingleby *et al.*, 2014; Walling *et al.*, 2014; Welch *et al.*, 2014). Here, selection is nonlinear, comprising a combination of stabilizing and disruptive processes as described by the fitness surfaces. These forms of selection coupled with the lack of net linear selection may contribute to the maintenance of genetic variation and prevent male sex combs from evolving to a single optimal phenotype. Furthermore, data from other systems suggest that disruptive selection, as we find here, may be important for divergent evolution and speciation (Schluter, 2000).

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